

All for One? Family Size and Children's Educational Distribution under Credit Constraints[†]

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Economists have been examining parents' decisions about their offspring's education in the context of "quantity-quality tradeoff" (e.g., Angrist, Lavy, and Schlosser 2010) or "birth order effects" (e.g., Black, Devereux, and Salvanes 2007). This article contributes to both literatures by examining the possibility that a child's schooling can be increased by having more siblings, instead of being diminished by competition for parents' resources, as has been suggested in most of the literature. Specifically, parents unable to finance their children's education may rely on some of their older children's labor income as a source of funding. Then the children selected for education will benefit from having more siblings who support them. This implies that the relationships among family size, education, birth order, and years of schooling may be systematically different across families depending on their access to assets and borrowing resources.

We first examine these relationships in a model combining convex returns to education and credit constraints. Our model predicts correlations among family size, years of schooling, and birth order, which would not exist when either of these two elements is absent. In particular, for credit constrained parents, our model predicts a positive correlation between family size and educational level of the most educated child. It also predicts that the schooling of a child will be positively correlated with his/her birth order as well

as with the fraction of female siblings (for males) when gender-biased norms exist. However, we show that these relationships weaken or disappear for parents with larger assets.

Second, we examine the model implications using datasets from the United States, Mexico, and South Korea. These countries are chosen not only because they have data on educational attainment of all adult children in a family but also because they are at different stages of development and have different cultural norms. We expect that the model implications generated by credit constraints will be more pronounced among low income families within a country and also more pronounced in low income countries (Mexico and Korea) than in more developed ones (United States). Gender-related predictions will also be more pronounced in Korea, which is known for its son preferences. We find the empirical results to be consistent with our model predictions. Although we do not rule out the possibility that an alternative model may generate some of the empirical patterns this article documents, it would be challenging to account for all systematic relationships between parents' income and intrafamily educational allocation using alternative hypotheses.

Apart from those cited above, our article is related to various studies. Tenikue and Verheyden (2010) examine cross-sibling transfers to understand children's schooling in Sub-Saharan Africa. Our article is different from theirs in that we examine the role of family size and birth order in schooling, while they focus on birth order effects within a family. We also allow for gender-based schooling choice and test the model implications in samples of siblings who completed their schooling and from countries that have very different socioeconomic settings. Morduch (2000) presents evidence that having more sisters can increase men's years of schooling in Africa, which can be explained by our model and empirical evidence. Behrman, Pollak

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and Taubman (1989) examine the possibility that a large family may face lower costs of education because of sibling-based policies. Their mechanism is distinct from ours because the former does not explain the differential impact we observe by father's education nor by birth order. Finally, our article contributes to the studies exploring intrafamily distribution of resources to support elderly parents (e.g., Raut and Tran 2005), in that this article suggests the same kind of intrafamily decisions could be made at an early stage when families decide children's schooling.

I. Model

Assume parents j with n_j number of children value their children's schooling as $\sum_{k=1}^{n_j} (h(S_{kj}) - (1 + c_{kj})S_{kj})$, where k denotes birth order (1 for the eldest child, and n_j for the youngest), S_{kj} denotes years of schooling of the k^{th} , c_{kj} is an idiosyncratic utility cost of schooling that is assumed independent of birth order.¹ Function h governs the returns to schooling, assumed such that $h(0) = 1$, $h' > 0$, and $h''(S) > 0$ for any $0 \leq S < a$ and $h'' < 0$ for any $S \geq a$.

Parents cannot borrow to finance the education of their children. The monetary cost of schooling, normalized as 1 per unit, must be financed out of parental assets A_j and labor earnings of siblings working as youth at a wage rate of w_k when they are not attaining the maximum education level $\bar{S}$. The wage rate is assumed to decrease in birth order.² Then the budget set is

$$\sum_{k=1}^{n_j} S_{kj} \leq A_j + \sum_{k=1}^{n_j} w_k(\bar{S} - S_{kj}).$$

Unconstrained parents will select S_{kj}^* such that $h'(S_{kj}^*) = 1 + c_{kj}$. When parents are credit constrained, the first- and second-order conditions ensure that for any two children k and k' ,

receiving a positive education, we must have $h''(S_{kj}), h''(S_{k'j}) < 0$,

$$(1) \quad \frac{h'(S_{kj}) - (1 + c_{kj})}{h'(S_{k'j}) - (1 + c_{k'j})} = \frac{1 + w_k}{1 + w_{k'}},$$

and that the monetary cost of their education is equal to resources:

$$\sum_{k=1}^{n_j} S_{kj}(1 + w_k) = A_j + \bar{S} \sum_{k=1}^{n_j} w_k.$$

PROPOSITION 1: *Depending on asset level, parents' schooling decision follows one of the three regimes: (i) For assets above $A^*(n_j)$, parents are unconstrained and select S_{kj}^* such that $h'(S_{kj}^*) = 1 + c_{kj}$. (ii) For parents with assets below $A^*(n_j)$ but above $\tilde{A}(n_j)$, all children receive a positive investment driven by the first-order conditions (1). (iii) For the rest of parents, at least one child acquires no schooling.*

PROOF:

This first regime occurs when the restriction is not binding, which occurs when assets are at least as large as $A^*(n_j) = \sum_{k=1}^{n_j} S_{kj}^* - w_k(\bar{S} - S_{kj}^*)$. Define $\tilde{A}(n_j)$ as being the asset level at which parents are indifferent between educating all children or spending all their resources on $n - 1$ of them. Consider family j which is indifferent between educating $n_j - 1$ or n_j children when it has asset $\tilde{A}(n_j)$. A slightly higher level of assets will enable them to increase the schooling investments of their children, and given that all children who receive positive education level must, by the second-order condition, fall into the range where the returns to schooling are concave, it will always be optimal to continue educating all children. For parental assets below $\tilde{A}(n_j)$, it can also be easily shown that even pooling all resources into a subset of children may not be sufficient to offer $S > b$ to these children, thus making it clearly optimal to invest in fewer children as assets decrease.

Now we examine the role of family size, birth order, and gender in our model.

Family Size.—In families with assets lower than $\tilde{A}(n_j)$, the most educated child in a family will benefit from having more siblings since their labor earnings will finance her schooling.

¹ While there is evidence that IQ and birth order are positively correlated, this does not appear to be linked to innate ability but more to posterior investments made by parents. See Black, Devereux, and Salvanes (2007).

² Since older children have more years where they can work to finance their younger sibling's education than vice versa.

However, for parents with assets between $\tilde{A}(n_j)$ and $A^*(n_j)$, having an additional sibling will penalize the most educated child as more siblings imply more competition for scarce resources. Finally, family size will be irrelevant for non-credit-constrained parents.

Birth Order.—In wealthy families, investment is independent of birth order since it depends only on the idiosyncratic cost of education. However, for credit-constrained parents, higher birth order children (i.e., younger ones) will be favored for schooling because wages are assumed to decrease in birth order (see equation (1) for the relationship between schooling and wage).

Gender.—Suppose that the opportunity cost of educating a girl is higher than that of a boy. For example, this condition would be satisfied in our model if females can earn a sufficiently higher wage when young or have to pay a higher utility cost to attend school. If parents are unconstrained, having an additional female sibling instead of a male will have no impact on that child's educational investment. In the case where parents are borrowing constrained but still are able to offer a positive educational investment in all children, having a female sibling will relax the parental borrowing constraint and offer male siblings more resources for their education. This pattern will be even more marked at lower levels of parental assets.

II. Empirical Strategy and Data

We empirically test the relationship between family size (n_j) and the years of schooling of the most educated child in a family (y_j) as follows:

$$y_j = \alpha n_j + \beta A_j + \gamma n_j \times A_j + \rho X_j + \varepsilon_j,$$

where A_j is proxied with the father's educational attainment. Our model suggests that $\alpha > 0$ but $\gamma < 0$; that is, for low income families, family size has positive impact on the most educated child's schooling, but that impact decreases as A_j increases. Additional control variables (X_j) include the cohort of the first child born interacted with A_j to capture differential schooling trends over time by father's educational level.

Next we use the following equation to examine the birth order effect:

$$(2) \quad y_{kj} = \alpha k + \beta A_j + \gamma k \times A_j \\ + \rho X_{kj} + \varepsilon_{kj},$$

where y_{kj} measures the educational attainment of the k^{th} birth order child in family j , and X_{kj} includes a gender dummy in addition to the controls mentioned previously. Our model suggests that $\alpha > 0$ but $\gamma < 0$. Note that standard errors are clustered at family level. Finally, to study the gender effect, we examine not only birth order but also siblings' gender composition within a similar framework.

We empirically test our model implications using the national surveys of the elderly in the United States, Korea, and Mexico. We use the "1993 Asset and Health Dynamics Among the Oldest Old" for the United States, the "2006 Korean Longitudinal Study of Aging" for Korea, and the "2001 Mexican Health and Aging Study" for Mexico, as they provide detailed information on the elderly respondents and their adult children. We use a subsample of families with more than or equal to two children to focus on educational choices across multiple children. We restrict the sample to fathers aged over 50 in the case of Korea and Mexico so that the children are likely to have completed their education (the US data included only fathers aged 70 and over).

Sample statistics show that the United States has the most educated fathers (11 years of schooling), followed by Korea (8 years), and then Mexico (5 years). However, in terms of children's education and family size, the United States and Korea are comparable to each other (13.5 years and 3.5 children), but Mexico on average has 9.5 years of schooling for children and 6 children per household. Using the difference between the average educational attainment of girls versus boys as an indication of gender-specific opportunity cost, we find that parents face higher costs of educating girls in Korea, followed by Mexico and then the United States.

III. Results

Table 1 shows the relationship between family size and the maximum educational attainment of

TABLE 1—FAMILY SIZE AND EDUCATION LEVEL OF THE MOST EDUCATED CHILD

	Korea	Mexico	US
n_j	0.372*** (0.050)	0.239*** (0.042)	-0.033 (0.168)
$n_j \times \text{primary}$	-0.190*** (0.069)	-0.316*** (0.052)	0.052 (0.174)
$n_j \times \text{secondary}$	-0.262*** (0.072)	-0.382*** (0.092)	0.046 (0.171)
$n_j \times \text{tertiary}$	-0.347*** (0.123)	-0.228* (0.131)	0.142 (0.175)
Observations	4,494	4,894	2,218

Notes: All regressions include dummies for father's educational attainment and the year the first child was born interacted with the educational category of the father.

***Significant at the 1 percent level.

**Significant at the 5 percent level.

*Significant at the 10 percent level.

the children within a family. For our analysis, we classify fathers into four categories based on their highest degree: no degree from formal schooling (baseline group), primary, secondary, and tertiary. In Mexico and Korea, having more siblings is positively correlated with the years of schooling of the most educated child when their fathers have no formal degree, but that relationship weakens as the father's education increases. The US data shows no significant relationship, which is to be expected because it is more developed than the other two countries, and, thus, credit constraints are less marked. Note that a positive correlation between family size and maximum years of schooling within a family can be the product of a simple statistical property where outliers are more likely in larger samples. However, this alternative explanation cannot justify the difference across fathers' educational attainments that we document, nor why we do not observe this pattern in the United States.

Table 2 presents the correlation between birth order and a child's years of schooling. In all countries, fathers with no formal education educate a child more than his/her adjacent older sibling by 0.2 to 0.7 years. This birth order effect is reduced for fathers with more education. For fathers with tertiary education, the effect remains negative only in the case of Korea, while in Mexico and the United States, it becomes positive, as has been shown for Nordic countries (Black, Devereux, and Salvanes 2005).

TABLE 2—BIRTH ORDER (k) AND YEARS OF EDUCATION

	Korea	Mexico	US
k_j	0.741*** (0.032)	0.293*** (0.019)	0.186** (0.078)
$k_j \times \text{primary}$	-0.265*** (0.043)	-0.235*** (0.026)	-0.142 (0.190)
$k_j \times \text{secondary}$	-0.456*** (0.044)	-0.525*** (0.044)	-0.231*** (0.082)
$k_j \times \text{tertiary}$	-0.519*** (0.053)	-0.767*** (0.068)	-0.294*** (0.083)
Observations	15,778	26,861	7,926

Notes: Standard errors clustered at the level of the family. Note of Table 1 applies.

A similar pattern is observed in terms of whether a child is the most educated in his family.

Next, we test whether in contexts where the opportunity cost of educating a girl is higher than that of a boy, a male from a poor family will benefit from having more sisters than brothers since these sisters would be likely to contribute to their education. We use a strategy similar to Vogl (2013): conditional on the number of younger siblings an individual has, the gender composition of the younger siblings is close to being random.³ Table 3 shows that men benefit from having a higher fraction of younger female siblings in Korea, and that this benefit decreases as fathers are more educated. We find no pattern in either Mexico or the United States, although the signs in the United States are consistent with the Korean example but much noisier. Note that we find females do not appear to suffer or benefit from having a different gender composition among their younger siblings.

Finally, while our model assumes that parents make all schooling decisions, it is likely that children will eventually also participate in this decision. Adding this to our model requires that children who sacrifice their education for the benefit of their siblings be compensated accordingly. While we do not observe transfers between siblings, we do measure transfers made from children to their elderly parents. Children who benefited from the investment of their

³ Our model would also suggest that the gender composition of older siblings would matter, but we do not use this variation in our regression analysis since that measure is likely to be biased in front of any gender-related stopping rule or sex-selective abortion.

TABLE 3—SIBLINGS' GENDER AND MALES' EDUCATION

	Korea	Mexico	US
Percent females	0.646** (0.295)	−0.358 (0.294)	1.502 (1.027)
Percent females × primary	−0.465 (0.354)	0.623* (0.346)	−0.991 (1.112)
Percent females × secondary	−0.740** (0.324)	0.577 (0.448)	−1.550 (1.041)
Percent females × tertiary	−0.467 (0.331)	0.649 (0.459)	−1.425 (1.049)
Observations	5,558	8,741	2,891

Notes: Standard errors clustered at the level of the family. Note of Table 1 applies.

siblings could thus compensate them by contributing more heavily to the care of their parents once they require assistance. We have explored this possibility in the case of Korea and find that a child who is from a poor family and receives the most education within his/her family has a much higher probability of making transfers to the parents, and the transfers are likely to be larger. Children from wealthy families do not appear to display this pattern.

IV. Conclusions

This paper develops a simple model where convex returns to education and credit constraints make siblings potential allies to obtain higher returns instead of competing for parents' investments. We also show that in a context where gender discrimination is high, female siblings may be particularly helpful to boys. The model implications are consistent with empirical patterns observed in the United States, Mexico, and Korea.

Our model has an important policy implication: in an environment where returns to education are sufficiently convex, a policy that restricts family size may lead to unintended adverse consequences. Suppose that the children selected for schooling compensate their siblings

for their sacrifice through transfers when they are grown up. Then a fertility restriction policy may impede this type of intrasibling arrangement among poor families and potentially harm all children, including those who would have worked for their siblings' schooling but get paid back later in life.

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